

Lecture 5

Matrix completion

Multi spike model with missing data.

parameters: rank $r \geq 2$, dimension $d \geq 2$, sample size $n \geq 2$, variance σ^2

unknown: $\theta^0 \in \mathbb{R}^{d \times d}$, $\text{rank}(\theta^0) \leq r$

observe: few missing entries of θ^0

$$(\underline{X}_1, \underline{Y}_1), \dots, (\underline{X}_n, \underline{Y}_n)$$

$$\underline{Y}_i = \langle \underline{X}_i, \theta^0 \rangle + \underline{W}_i, \quad \underline{X}_i = e_{\underline{a}(i)} e_{\underline{b}(i)}^T$$

$$\underline{W}_1, \dots, \underline{W}_n \stackrel{\text{i.i.d.}}{\sim} N(0, \sigma^2), \quad \underline{a}(1), \dots, \underline{a}(n), \underline{b}(1), \dots, \underline{b}(n) \sim \text{Unif}(\{1, \dots, d\})$$

error measure: $\frac{1}{d^2} \|\hat{\theta} - \theta^0\|_F^2$

necessary: $n \gg r \cdot d$, # observation $\gg$ # dof

$n \geq d \cdot \log d$, all rows & columns covered

additional obstruction $r=1, \sigma=0$

Suppose $\theta^0 \in_r \{\pm M \cdot e_i e_i^T\}$

$$\mathbb{P}(\text{loc } 1,1 \text{ not observed}) = \left(1 - \frac{1}{d^2}\right)^n \geq 1 - \frac{n}{d^2} \rightarrow 1$$

$$\mathbb{E}\left[\frac{1}{d^2} \|\hat{\theta} - \theta^0\|_F^2 \mid \text{loc } 1,1 \text{ not observed}\right] \geq \frac{M^2}{d^2}$$

$\rightarrow$ will allow error to grow with $\|\theta^0\|_{\max}$

Theorem $\exists$ poly-time estimator $\hat{\theta}$, if $n \geq d \cdot \log d$, then w.h.p.

$$\frac{1}{d^2} \|\hat{\theta} - \theta^0\|_F^2 \leq \frac{(\sigma^2 + \|\theta^0\|_{\max}^2) \cdot r \cdot d \cdot \log(d)}{n}$$

Strategy for poly-time estimator

two steps

① Rough estimation (simple function of observation) $\mathbb{E}[Y] = \frac{1}{d^2} \sum_{\underline{a}, \underline{b}} \theta_{\underline{a}, \underline{b}}^0 \cdot d^2 \cdot e_{\underline{a}} e_{\underline{b}}^T = \theta^0$

② Boost error bound (reduce to our multi-spike matrix estimator / analysis)

Compute $\underline{Y}$ as function of $(\underline{X}_1, \underline{Y}_1), \dots, (\underline{X}_n, \underline{Y}_n)$ ensure $\mathbb{E}[\underline{Y}] = \theta^0$

$$\text{compute: } \underline{Y} = \frac{1}{n} (\underline{Y}_1 \underline{X}_1^T \cdot d^2 + \underline{Y}_2 \underline{X}_2^T \cdot d^2 + \dots + \underline{Y}_n \underline{X}_n^T \cdot d^2)$$

$$\text{output: } \hat{\theta} = \argmin \{\|\theta - \underline{Y}\|_F^2 \mid \text{rank}(\theta) \leq r\}$$

$$\underline{Y} = \theta^0 + \underline{W}, \quad \mathbb{E}[\underline{W}] = 0$$

* A noisy version of θ^0

error analysis

$$\text{let } \underline{U} = \hat{\theta} - \theta^0, \quad \underline{W} = \underline{Y} - \theta^0$$

$$\text{claim: } \|\underline{U}\|_F^2 \leq 8r \cdot \|\underline{W}\|^2$$

$$\begin{aligned} \underline{W} &= \underline{Y} - \theta^0 = \frac{1}{n} \sum_{i=1}^n (\underbrace{\underline{Y}_i d^2 \underline{X}_i^T}_{\underline{Z}_i^{(1)}} - \theta^0) \quad * \underline{Y}_i = \langle \underline{X}_i, \theta^0 \rangle + \underline{W}_i, \quad \underline{W}_i \sim N(0, \sigma^2) \\ &\quad \underbrace{(\langle \underline{X}_i, \theta^0 \rangle d^2 \underline{X}_i^T)}_{\underline{Z}_i^{(1)}} + \underbrace{\underline{W}_i \cdot d^2 \underline{X}_i^T}_{\underline{Z}_i^{(2)}} \\ &= \frac{1}{n} \sum_{i=1}^n \underline{Z}_i^{(1)} + \frac{1}{n} \sum_{i=1}^n \underline{Z}_i^{(2)} \end{aligned}$$

claim if $n \geq d \cdot \log d$, w.h.p.

- $\left\| \frac{1}{n} \sum_{i=1}^n \underline{z}_i^{(1)} \right\|^2 \leq \|\theta^0\|_{\max}^2 \cdot \frac{d^3 \log d}{n}$
- $\left\| \frac{1}{n} \sum_{i=1}^n \underline{z}_i^{(1)} \right\|^2 \leq \sigma^2 \cdot \frac{d^3 \log d}{n}$

Assuming the claim

$$\frac{1}{d^2} \|\underline{u}\|_F^2 \leq \frac{8r}{d^2} \|\underline{w}\|^2$$

claim $\lesssim \frac{r}{d^2} (\|\theta\|_{\max}^2 + \sigma^2) \frac{d^3 \log d}{n}$

Matrix Bernstein bound

$\underline{z}$ random $d \times d$ matrix, $d \geq 2$

$$\mathbb{E}[\underline{z}] = 0, \{\underline{z}\} = \{\underline{z}^T\}$$

(σ, b)-Bernstein condition

$$\forall j \in \mathbb{N}_{\geq 2}, \|\mathbb{E}[\|\underline{z}\|^{j-2}] \underline{z} \underline{z}^T\| \leq j! \cdot b^{j-2} \cdot \sigma^2$$

Theorem If $\underline{z}$ satisfies (σ, b)-Bernstein condition, then With prob. $\geq 1 - d^{-100}$

$$\left\| \frac{1}{n} \sum_i \underline{z}_i \right\| \leq \sigma \sqrt{\frac{\log d}{n}} + b \frac{\log d}{n}$$

Loewner order A, B symmetric matrices

$$A \preceq B \Leftrightarrow B - A \text{ is p.s.d.}$$

$$\Leftrightarrow \forall v, \langle v, Av \rangle \leq \langle v, Bv \rangle$$

fact if $-A \leq B \leq A$, then $\|B\| \leq \|A\|$

$$\underline{z}^{(1)} = \langle \underline{x}_i, \theta^0 \rangle \cdot d^2 \cdot \underline{x} - \theta^0, \quad \underline{x} = e_a \cdot e_b^T$$

claim $\underline{z}^{(1)}$ is (σ, b) -Bernstein for $\sigma \leq \|\theta^0\|_{\max} \cdot d^{3/2}$ and $b \leq \|\theta^0\|_{\max} \cdot d^2$

Proof

$$\|\underline{z}^{(1)}\| \leq \underbrace{\|\langle \underline{x}, \theta^0 \rangle \cdot d^2 \cdot \underline{x}\|}_{\|\theta^0\|_{\max} \cdot d^2} + \|\theta^0\| \leq \|\theta^0\|_{\max} \cdot d^2$$

p.s.d.

$$\begin{aligned} 0 &\leq \underbrace{\mathbb{E}[\underline{z}^{(1)}(\underline{z}^{(1)})^T]}_{\text{p.s.d.}} = \mathbb{E}[\langle \underline{x}, \theta^0 \rangle^2 \cdot d^4 \underline{x} \underline{x}^T + \theta^0 (\theta^0)^T - 2\theta^0 (\theta^0)^T] \quad * \mathbb{E}[(X - \mathbb{E}[X])^2] \\ &= \mathbb{E}[\langle \underline{x}, \theta^0 \rangle^2 \cdot d^4 \underline{x} \underline{x}^T - \theta^0 (\theta^0)^T] \quad = \mathbb{E}[X^2] - (\mathbb{E}[X])^2 \\ &= \mathbb{E}[(\theta_a^0 e_b)^2 \cdot d^4 \cdot e_a e_a^T] \quad e_a e_a^T (e_a e_b^T)^T = e_a (e_b^T e_a) e_a^T = e_a e_a^T \\ &\leq \|\theta^0\|_{\max}^2 \cdot d^4 \cdot \mathbb{E}[e_a e_a^T] = \|\theta^0\|_{\max}^2 d^3 \cdot \mathbb{I} \\ \|\mathbb{E}[\underline{z} \underline{z}^T]\| &\leq \|\theta^0\|_{\max}^2 \cdot d^3 \end{aligned}$$

$$\| \mathbb{E}[\| \underline{z}^{(j)} \|^{j-2} \cdot \underline{z}^{(j)} (\underline{z}^{(j)})^T] \| \leq (2 \|\theta^0\|_{\max} \cdot d^{\frac{1}{2}})^{j-2} \cdot \underbrace{\| \mathbb{E}[\underline{z}^{(j)} (\underline{z}^{(j)})^T] \|}_{\underbrace{\|\theta^0\|_{\max}^2 \cdot d^3}_{\sigma_1^2}}$$

Theorem

$$\begin{aligned} \text{w.h.p.}, \quad \| \frac{1}{n} \sum_{i=1}^n \underline{z}_i \| &\leq \sigma \sqrt{\frac{\log d}{n}} + b \frac{\log d}{n} \\ &\leq \|\theta^0\|_{\max} \cdot \left(\frac{d^{\frac{3}{2}} \sqrt{\log d}}{\sqrt{n}} + \frac{d^2 \log d}{n} \right) \end{aligned}$$

$$\underline{z}^{(j)} = \underline{w} \cdot d^{\frac{j}{2}} \cdot \underline{x}, \quad \underline{w} \sim N(0, \sigma^2)$$

claim $\underline{z}^{(j)}$ is (σ_2, b_2) -Bernstein, for $\sigma_2 =$, $b_2 =$

Proof $\| \underline{z}^{(j)} \| = |\underline{w}| \cdot d^{\frac{j}{2}}$

$$\mathbb{E}[\| \underline{z}^{(j)} \|^{j-2} \cdot \underline{z}^{(j)} (\underline{z}^{(j)})^T] = \underbrace{\mathbb{E}[|\underline{w}|^j] \cdot d^{\frac{3j}{2}} \cdot \frac{1}{d} \text{Id}}_{|\underline{w}|^j \cdot (d^{\frac{1}{2}})^j \cdot \underline{e}_0 \underline{e}_0^T} \leq j! \underbrace{(\sigma d^{\frac{1}{2}})^{j-2}}_{b_2} \cdot \underbrace{\sigma^2 \cdot d^3}_{\sigma_2^2}$$